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Minimax Lower Bound for Low-Rank Matrix-Variate Logistic Regression

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Outline

- Motivation and Model Overview
 - Theoretical Result
 - Construction of Our Theory
- } Part 1
- } Part 2

Matrix-Variate Logistic Regression: Probability Model

Vector Logistic Regression:

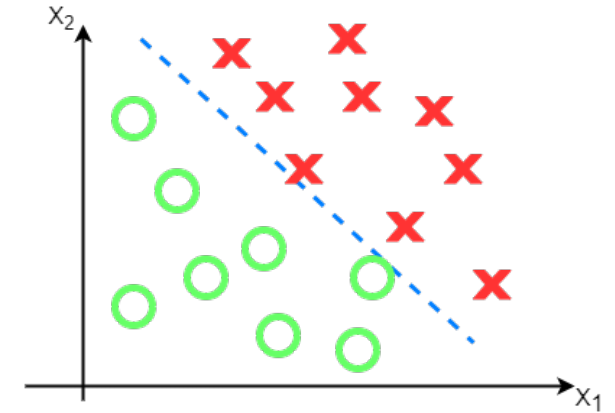
$$\mathbb{P}_{y|\mathbf{x}}(y_i = 1|\mathbf{x}_i) = \frac{1}{1 + \exp(-(\mathbf{b}^T \mathbf{x}_i + z))}$$

$y_i \in \{0, 1\}$: binary response (output class)

$\mathbf{b} \in \mathbb{R}^m$: unknown coefficient vector

$\mathbf{x}_i \in \mathbb{R}^m$: covariate (data sample)

z : zero-mean intercept (bias)



Matrix Logistic Regression:

$$\mathbb{P}_{y|\mathbf{X}}(y_i = 1|\mathbf{X}_i) = \frac{1}{1 + \exp(-(\langle \mathbf{B}, \mathbf{X}_i \rangle + z))}$$

$\mathbf{B} \in \mathbb{R}^{m_1 \times m_2}$

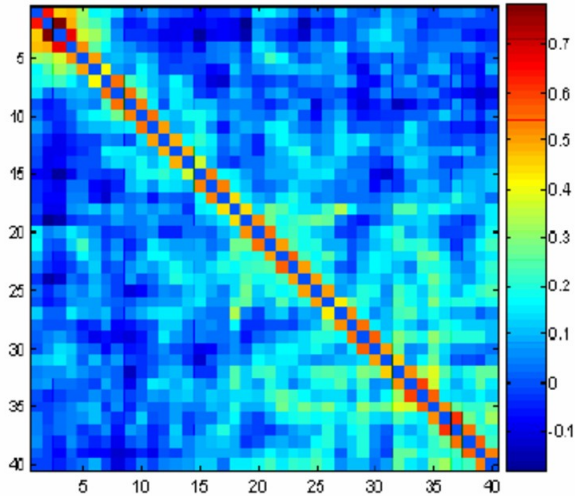
$\mathbf{X}_i \in \mathbb{R}^{m_1 \times m_2}$

Due to the inner product, both models are mathematically equivalent.

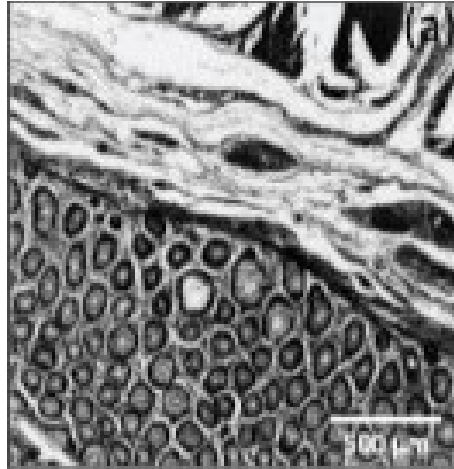
Why Matrix-Variate Logistic Regression?

$$\mathbb{P}_{y|\mathbf{X}}(y_i = 1|\mathbf{X}_i) = \frac{1}{1 + \exp(-(\langle \mathbf{B}, \mathbf{X}_i \rangle + z))}$$

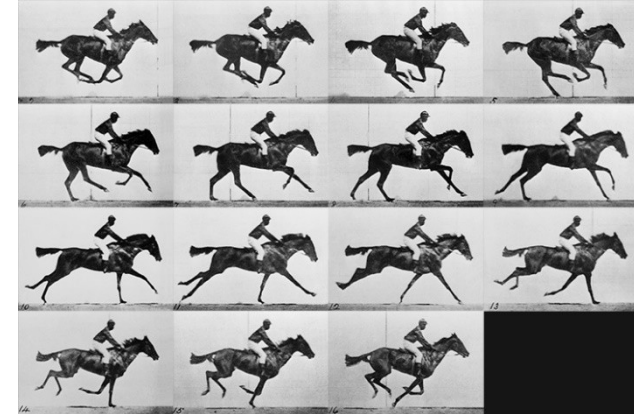
- In many practical applications covariates naturally take the form of two-dimensional arrays, such as:



Electroencephalography
(EEG) data



Fiber-bundle Imaging



Spatial-temporal data

- The coefficients are also matrices, and contain rich information in their spatial structure.

Why Matrix-Variate Logistic Regression?

$$\mathbb{P}_{y|\mathbf{X}}(y_i = 1|\mathbf{X}_i) = \frac{1}{1 + \exp(-(\langle \mathbf{B}, \mathbf{X}_i \rangle + z))}$$

- For estimating \mathbf{B} , classical machine learning techniques vectorize the data and estimate a coefficient vector.



Low-Rank Matrix-Variate Logistic Regression

$$\mathbf{b} \in \mathbb{R}^{m_1 m_2}$$

Why Low-Rank Matrix-Variate Logistic Regression?

$$\mathbb{P}_{y|\mathbf{x}}(y_i = 1|\mathbf{X}_i) = \frac{1}{1 + \exp(-(\langle \mathbf{B}, \mathbf{X}_i \rangle + z))}$$

- Low-rank structures may arise from the presence of redundant variables.
- The model's intrinsic degrees of freedom are smaller than its extrinsic dimensionality.

We can represent the data in a lower dimensional space

We can reduce the sample complexity of estimating the parameters

Prior Work:

- Vector based logistic regression
 - High-dimensional logistic regression [e.g F. Abramovich and V. Grinshtein 2018]
- Regularized matrix-variate logistic regression
 - Regularization for rank-optimized or sparse coefficient estimation [e.g J. Zhang and J. Jiang 2018, J. V. Shi et al 2014]
 - Regularization for inference on image data [e.g B. An and B. Zhang 2020]

Minimax Lower Bounds Provide Error Thresholds

Why Minimax Lower Bounds?

- They provide insights to:
 - The fundamental error thresholds of the estimation problem and the performance of corresponding algorithms.
- Indicate the parameters on which the minimax risk depends.

Prior Work

- Minimax lower bounds for graph-based logistic regression [e.g Q. Berthet and N. Baldin 2020].

Outline of This Work

- Derive a minimax lower bound that is proportional to the rank and dimensions of the coefficient matrix.
- Reduce the sample complexity from the vector setting.
- Show that the methods used are easily extendible to the tensor case.

Model and Problem Formulation

- Consider the matrix LR problem:

$$\mathbb{P}_{y|\mathbf{x}}(y_i = 1|\mathbf{X}_i) = \frac{1}{1 + \exp(-(\langle \mathbf{B}, \mathbf{X}_i \rangle + z))}$$

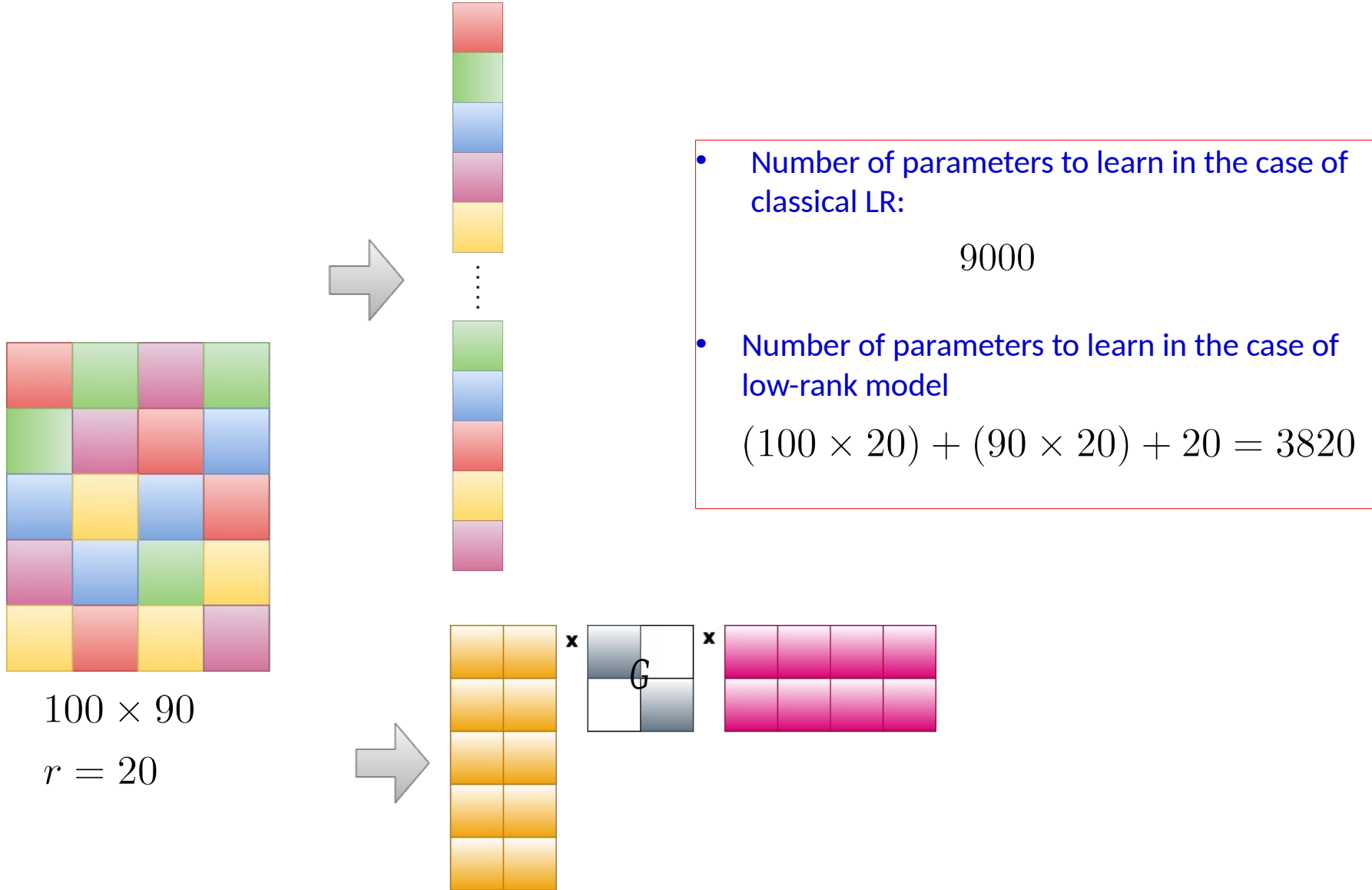
- Goal: Find estimate $\hat{\mathbf{B}}$ of \mathbf{B} using training data $\{\mathbf{X}_i, y_i\}_{i=1}^n$.
- Consider the case where \mathbf{B} is a rank- r matrix. Specifically, the rank- r singular value decomposition of \mathbf{B} is

$$\mathbf{B} = \mathbf{B}_1 \mathbf{G} \mathbf{B}_2^T \quad \begin{bmatrix} | & & | \\ \mathbf{b}_1^1 & \cdots & \mathbf{b}_1^r \\ | & & | \end{bmatrix} \begin{bmatrix} \lambda_1 & & \\ & \ddots & \\ & & \lambda_r \end{bmatrix} \begin{bmatrix} - & \mathbf{b}_2^1 & - \\ & \vdots & \\ - & \mathbf{b}_2^r & - \end{bmatrix}$$

$\mathbf{B}_1 \in \mathbb{R}^{m_1 \times r}$ } Matrix of left/right singular vectors
 $\mathbf{B}_2 \in \mathbb{R}^{m_2 \times r}$ } (with orthonormal columns) $\mathbf{G} = \text{diag}(\lambda_1, \dots, \lambda_r) \in \mathbb{R}^{r \times r}$, $\lambda_i > 0 \forall i \in [r]$ } Matrix of singular values

$$\mathbb{P}_{y|\mathbf{x}}(y_i = 1|\mathbf{x}_i) = \frac{1}{1 + \exp(-(\langle \mathbf{B}_1 \mathbf{G} \mathbf{B}_2^T, \mathbf{X}_i \rangle + z))}$$

Benefit of Low-Rank Models (Toy Example)



Model and Problem Formulation, Continued

Consider the parameter space, \mathcal{P}_r , of all rank- r matrices in $\mathbb{R}^{m_1 \times m_2}$, and a subset $\mathcal{B}_d \subset \mathcal{P}_r$ of rank- r matrices with finite energy. More formally,

$$\mathcal{B}_d(\mathbf{0}) \triangleq \{\mathbf{B}' \in \mathcal{P}_r : \|\mathbf{B}' - \mathbf{0}\|_F < d\}$$

The minimax risk is thus defined as the worst-case mean squared error (MSE) for the best estimator, i.e.,

$$\varepsilon^* = \inf_{\hat{\mathbf{B}}} \sup_{\mathbf{B} \in \mathcal{B}_d(\mathbf{0})} \mathbb{E}_{\mathbf{y}, \underline{X}^c} \left\{ \|\hat{\mathbf{B}} - \mathbf{B}\|_F^2 \right\}$$

Main Results of Prior Works

Vector-based Logistic Regression:

$$\mathcal{O} = \frac{m_1 m_2}{n}$$

Matrix Logistic Regression: ??

Main Result

Theorem 1 [Taki et al. 2021]

Consider the rank- r matrix LR problem with n i.i.d observations, $\{\mathbf{X}_i, y_i\}_{i=1}^n$ where the true coefficient matrix $\|\mathbf{B}\|_F^2 < d^2$.

Then, for covariate $\text{vec}(\mathbf{X}_i) \sim \mathcal{N}(0, \sigma^2 \mathbf{I}_{m_1 m_2})$ the minimax risk is lower bounded by

$$\varepsilon^* \geq \frac{\left(\left[c_2 (c_1 \mathbf{r} (\mathbf{m}_1 + \mathbf{m}_2 - \mathbf{2}) + c_1 (\mathbf{r} - \mathbf{1})) - c_3 \right] - 1 \right)}{8n\sigma \sqrt{\frac{2}{\pi}}}$$

where

$$c_1 = \left(1 - \frac{1}{10}\right)^2, \quad c_2 = \frac{\log_2(e)(\sqrt{2} - 1)}{4\sqrt{2}}, \quad c_3 = \left(\frac{3(\sqrt{2} - 1)}{\sqrt{8}}\right) \log_2\left(\frac{3}{2}\right)$$

Main Result and Discussion

$$\epsilon^* \geq \frac{\left(\left[c_2 (c_1 \mathbf{r} (\mathbf{m}_1 + \mathbf{m}_2 - \mathbf{2}) + c_1 (\mathbf{r} - \mathbf{1})) - c_3 \right] - 1 \right)}{8 \mathbf{n} \sigma \sqrt{\frac{2}{\pi}}}$$

- Compared to the vector case, result shows a decrease in the lower bound.

Minimax risk for vector based LR:

$$\mathcal{O}\left(\frac{m_1 m_2}{n}\right)$$

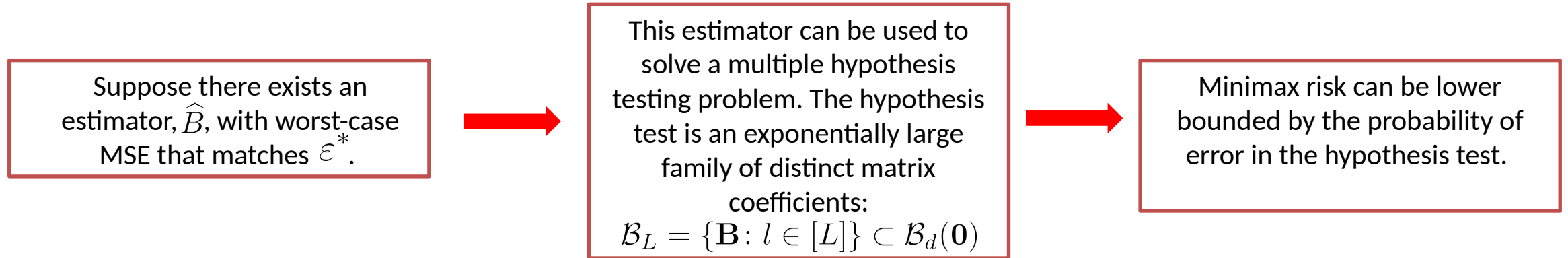
Minimax risk for rank-r matrix LR:

$$\mathcal{O}\left(\frac{r(m_1 + m_2 + 1)}{n}\right)$$

- Lower bound on the minimax risk is proportional to the intrinsic degrees of freedom in the coefficient matrix LR.

The Exciting Part! Proof of Main Results

Proof of Theorem 1 uses an argument based on Fano's inequality, more specifically:



Our goal: Further lower bound the probability of error.

Action items:

- Construct \mathcal{B}_L
- Find upper and lower bounds on the conditional mutual information $\mathbb{I}(\mathbf{y}; l | \underline{\mathbf{X}}^c)$

The Exciting Part! Proof of Main Results

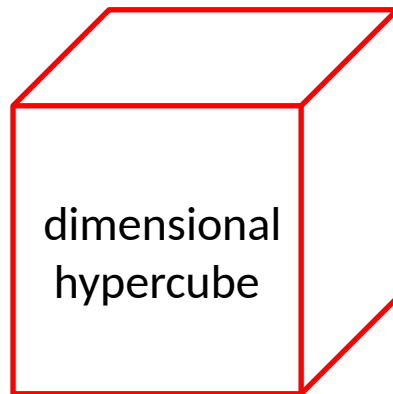
1. Constructing \mathcal{B}_L

a) We must **construct** \mathcal{B}_L such that a **minimum distance condition holds**, namely:

$$\|\mathbf{B}_l - \mathbf{B}_{l'}\|_F^2 \geq 8\delta$$

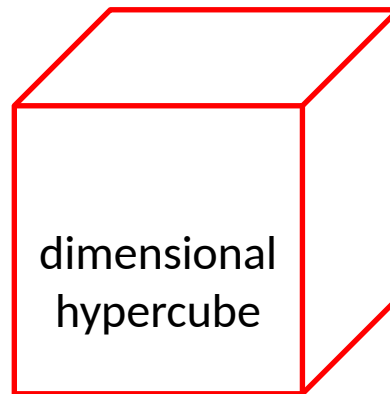
b) Since $\mathbf{B}_l = \mathbf{B}_1 \mathbf{G} \mathbf{B}_2^T$, we must **construct three separate sets** and **derive conditions** under which they exist simultaneously

Hypercube method: Construct a set of binary vectors/matrices with a minimum distance between any two distinct elements



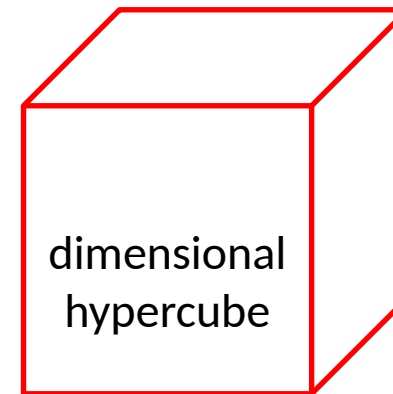
$$\mathbf{G}_f | f \in [F]$$

- Square diagonal matrix



$$\mathbf{B}_{1,p_1} | p_1 \in [P_1]$$

- Orthonormal Columns
- Bounded energy



$$\mathbf{B}_{1,p_2} | p_2 \in [P_2]$$

- Orthonormal Columns
- Bounded energy

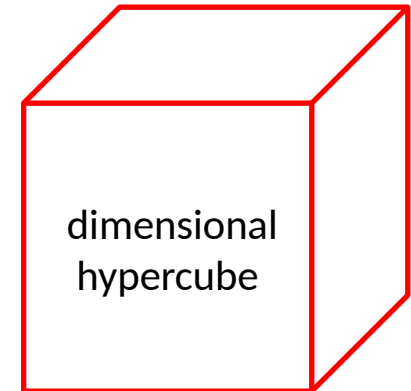
The Exciting Part! Proof of Main Results

Lemma 1: “Each hypercube exists with probability”

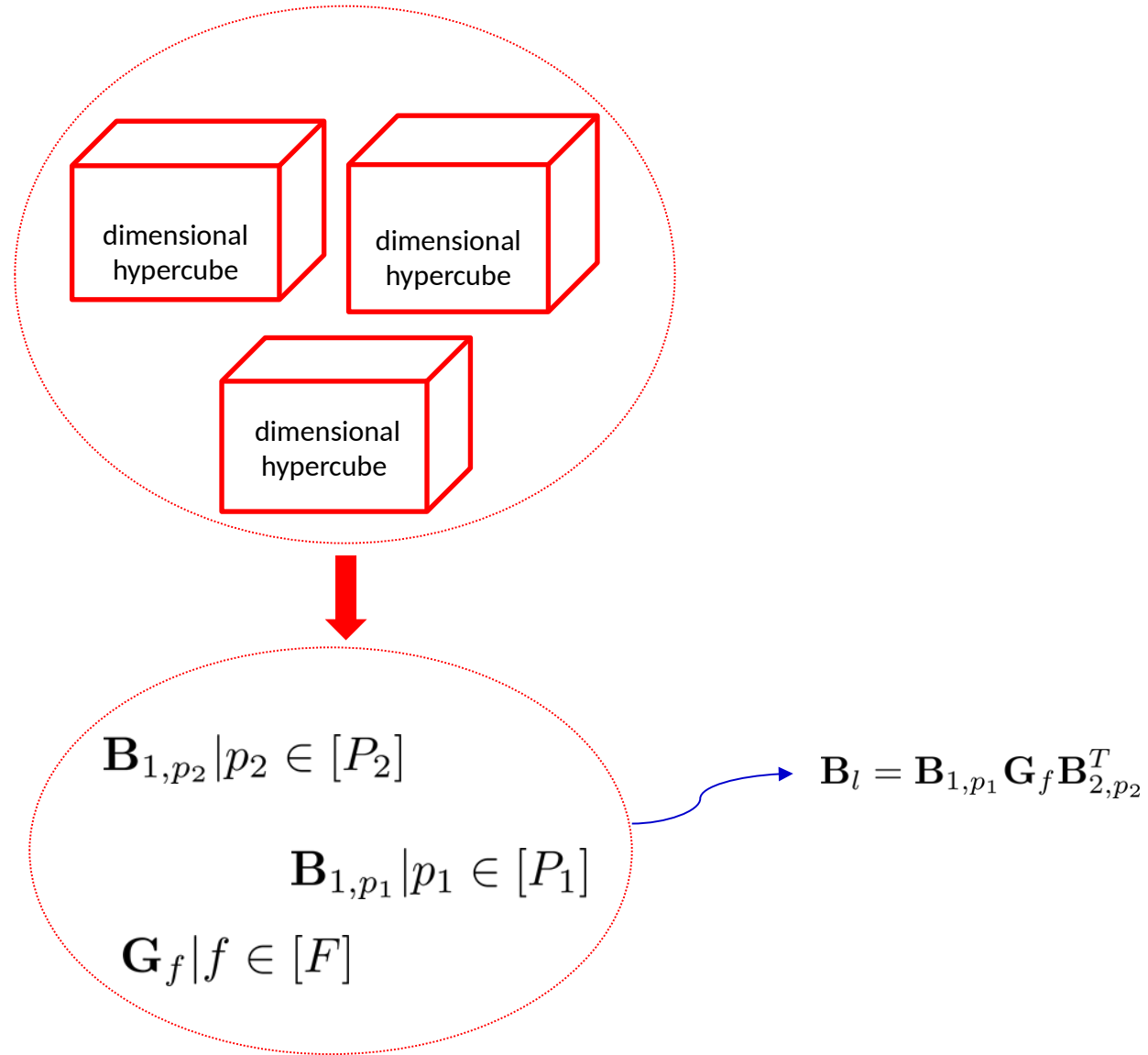
Lemma 1

Let $r > 0$ and $F \geq 2$. Consider the set of F vectors $\{\mathbf{s}_f \in \mathbb{R}^{r-1} : f \in [F]\}$, where each entry in vector \mathbf{s}_f is an independent and identically distributed random variable taking values $\left\{-\frac{1}{\sqrt{r-1}}, +\frac{1}{\sqrt{r-1}}\right\}$ uniformly. The probability that there exists a distinct pair (f, f') such that $\|\mathbf{s}_f - \mathbf{s}_{f'}\|_0 < \frac{r-1}{20}$ is upper bounded as follows:

$$\begin{aligned} & \mathbb{P}(\exists (f, f') \in [F] \times [F], f \neq f' : \|\mathbf{s}_f - \mathbf{s}_{f'}\|_0 < \frac{r-1}{20}) \\ & \leq \exp \left[2 \log(F) - \log(2) - \frac{1}{2} \left(1 - \frac{1}{10}\right)^2 (r-1) \right]. \end{aligned} \quad (1)$$



The Exciting Part! Proof of Main Results



The Exciting Part! Proof of Main Results

Lemma 2: “For all sets to exists simultaneously, we can construct set \mathcal{B}_L with L elements, where the distance between any two elements is bounded”

Lemma 2

There exists a collection of L matrices $B_L \triangleq \{\mathbf{B}_l : l \in [L]\} \subset \mathcal{B}_d(\mathbf{0})$ for some $d > 0$ of cardinality

$$L = 2^{\lfloor \frac{\log_2(e)}{4} \left(\left(1 - \frac{1}{10}\right)^2 (r(m_1 + m_2 - 1)) + \left(1 - \frac{1}{10}\right)^2 (r - 1) \right) - \frac{3}{2} \log_2\left(\frac{3}{2}\right) \rfloor} \quad (1)$$

such that for any

$$\sqrt{\frac{8(r-1)}{r}} < \varepsilon \leq d \sqrt{\frac{r-1}{r}}, \quad (2)$$

we have

$$\frac{r\varepsilon^2}{r-1} < \|\mathbf{B}_l - \mathbf{B}_{l'}\|_F^2 \leq 4 \frac{r\varepsilon^2}{r-1}. \quad (3)$$

Our packing:

The Exciting Part! Proof of Main Results

1. Bounding $\mathbb{I}(\mathbf{y}; l | \underline{\mathbf{X}}^c)$

a) Lower bound using Fano's inequality

- We require the existence of an estimator producing estimate $\hat{\mathbf{B}}$ and achieving minimax lower bound $\varepsilon^* = \sqrt{\delta}$
- Consider the minimum distance decoder: $\hat{l}(\mathbf{y}) \triangleq \arg \min_{\mathbf{B}_{l'} \in \mathcal{B}_d(\mathbf{0})} \|\hat{\mathbf{B}} - \mathbf{B}_{l'}\|_F^2$

$$\|\hat{\mathbf{B}} - \mathbf{B}_l\|_F^2 < \sqrt{2\delta}: \text{detect } \mathbf{B}_l \text{ and } \mathbb{P}(\hat{l}(\mathbf{y}) \neq l) = 0$$

$$\|\hat{\mathbf{B}} - \mathbf{B}_l\|_F^2 \geq \sqrt{2\delta}: \text{detection error might occur}$$

$$\mathbb{P}(\hat{l}(\mathbf{y}) \neq l) \leq \mathbb{P}\left(\|\hat{\mathbf{B}} - \mathbf{B}_l\|_F^2 \geq \sqrt{2\delta}\right)$$

$$\text{Fano's inequality states that: } \mathbb{I}(\mathbf{y}; l) \geq \left(1 - \mathbb{P}(\hat{l}(\mathbf{y}) \neq l)\right) \log_2(L) - 1 \triangleq u_1$$

The Exciting Part! Proof of Main Results

1. Bounding $\mathbb{I}(\mathbf{y}; l | \underline{\mathbf{X}}^c)$

b) Upper bound using

$$\mathbb{I}(\mathbf{y}; l | \underline{\mathbf{X}}^c) \leq \frac{1}{L^2} \sum_{l, l'} \mathbb{E}_{\underline{\mathbf{X}}^c} D_{KL}(f_l(\mathbf{y} | \mathbf{X}) || f_{l'}(\mathbf{y} | \mathbf{X})) \triangleq u_2$$

Lemmas 3 and 4 provide upper and lower bounds:

$$\frac{\sqrt{2}-1}{\sqrt{2}} \log_2 L - 1 \leq \mathbb{I}(\mathbf{y}; l | \mathbf{X}) \leq n\sigma \frac{2}{r} \sqrt{\frac{2}{\pi}} \varepsilon.$$

Some Closing Remarks

The result is interesting because:

- The analysis is non-trivial because the model uses a logistic function. Moreover, the result explicitly leverages the low-rank structure thus the hypothesis set is constructed from three factor sets. We derive conditions under which all sets can exist, and can be generalized to the tensor case.
- Two **hypotheses may be far apart** but produce the **same model** (or same observation). Our result gives insight into the parameters in which an achievable minimax risk might depend.

Current Investigations and Future Work

Study the benefits of imposing similar low-rank structures in the multi-dimensional LR setting:

Minimax risk lower bounds on the coefficient estimation in tensor-variate logistic regression.

Develop algorithms that meet the minimax lower bounds.

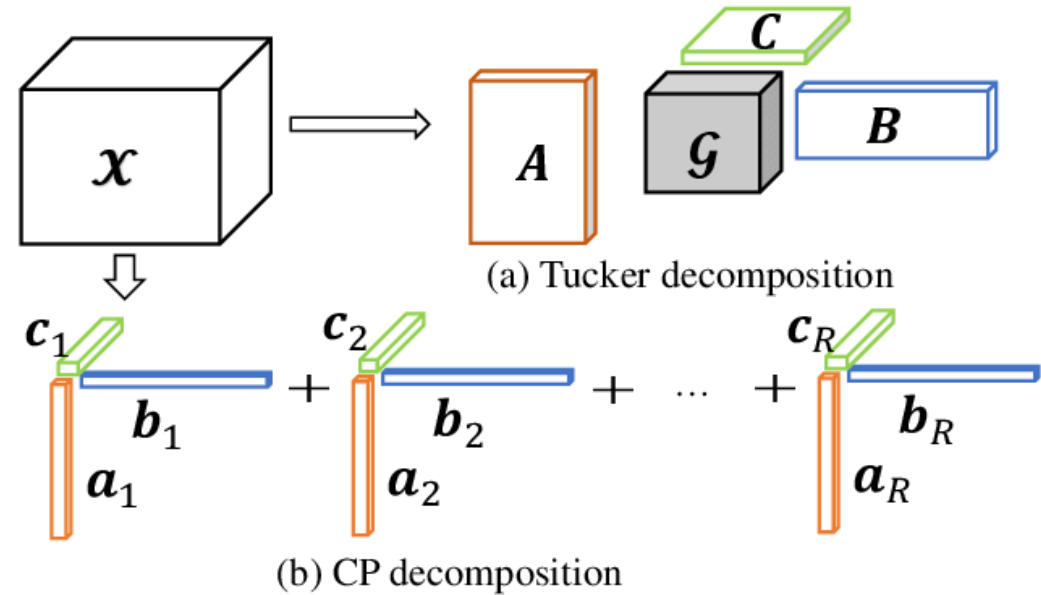
Test the performance of these algorithms on practical data.

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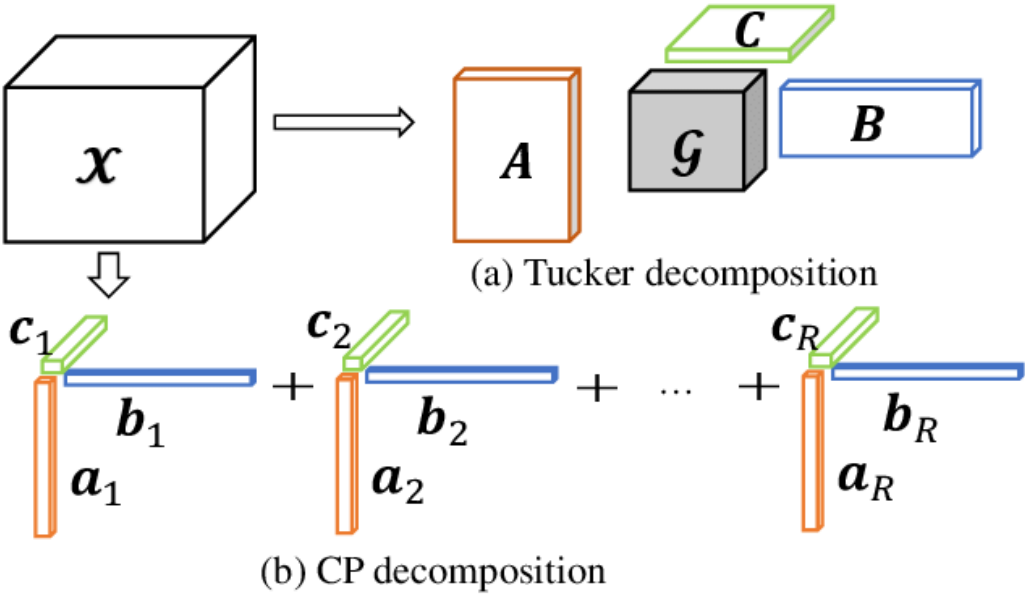
- CANDECOMP/PARAFAC (CP).
- Low-rank Tucker.



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